# British, American, and British-American Social Mobility: Intergenerational Occupational Change Among Migrants and Non-Migrants in the Late 19th Century

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# **RESEARCH QUESTIONS**

In previous work (Long & Ferrie 2011; Long & Ferrie 2007) we examine trends in intergenerational social mobility in the U.S. and Britain in the nineteenth century.

Key finding: mobility in U.S. in 19<sup>th</sup> century significantly greater than in Britain, unlike the present.

Along with theoretical results (Piketty 1995, Benabou and Ok 2001, Benabou and Tirole 2006) helps explain durability of "myth" of exceptional American mobility.

## **RESEARCH QUESTIONS**

The present study adds to our comparison of men in Britain and the U.S. the most (geographically) mobile group: trans-Atlantic migrants from Britain to the U.S.. We want to know

- How much intergenerational mobility did this group experience?
- How did their mobility experience compare with that of non-migrants in both countries?
- What can be said about the selectivity of the migrants?

#### **BACKGROUND**

The "quality" of immigrants is usually assessed by examining how they do relative to the native-born

But this cannot distinguish between change in overall home-country quality and change in the selectivity of immigration

Focuses exclusively on immigrants' experience after arrival in the destination

#### **BACKGROUND**

A complementary literature focuses on the "brain drain": selective immigration's impact on home-country characteristics

Focuses exclusively on migrants' experience before departure in the home country and the non-migrants' experience in the home country before and after migrants depart

#### **BACKGROUND**

Few studies examine the (1) migrants before departure from home and after arrival at destination and (2) non-migrants before and after the migrants depart (Abramtizky et al. 2010; Wegge 2002)

A different perspective on "selectivity"

But data on both "movers" & "stayers" is seldom available

#### **OUR APPROACH**

Here, we use 2 cohorts of British movers and stayers (1861-1880 & 1881-1900), observing (1) migrants before & after departure and (2) non-migrants before & after the migrants left

We account for selection explicitly

We provide the first measurement of intergenerational mobility for one of the largest groups of migrants to the U.S.

#### THE CONTEXT

Migration was completely unrestricted at this time (before the Quota System of the 1920s)

Driven not by desperation (c.f. Irish Famine migrants) but by "normal" forces (e.g. relative wages)

The British were a large fraction of the migrant stream (close to 40% in some years), but their share moved opposite the total volume of migration

#### THE CONTEXT

The Britain each cohort left behind was a decade or more ahead of the U.S. in its industrialization

More opportunity in the U.S. for those squeezed out by changes (consolidation in farming, displacement of craft workers by factories and machines)

Previously, we created samples of males linked across censuses from 1861-1881 & 1881-1901 in Britain, and males linked from 1860-1880 & 1880-1900 in the U.S.

Linkage based on (i) name, (ii) year of birth, (iii) parish & county (Britain) or state (U.S.) of birth.

Individuals were 30-39 years old in the terminal year and were observed with their fathers in the initial year.

Fathers' & sons' occupations observed at same lifecycle point .

For comparable data on migrants from Britain to the U.S., we generated 2 new samples

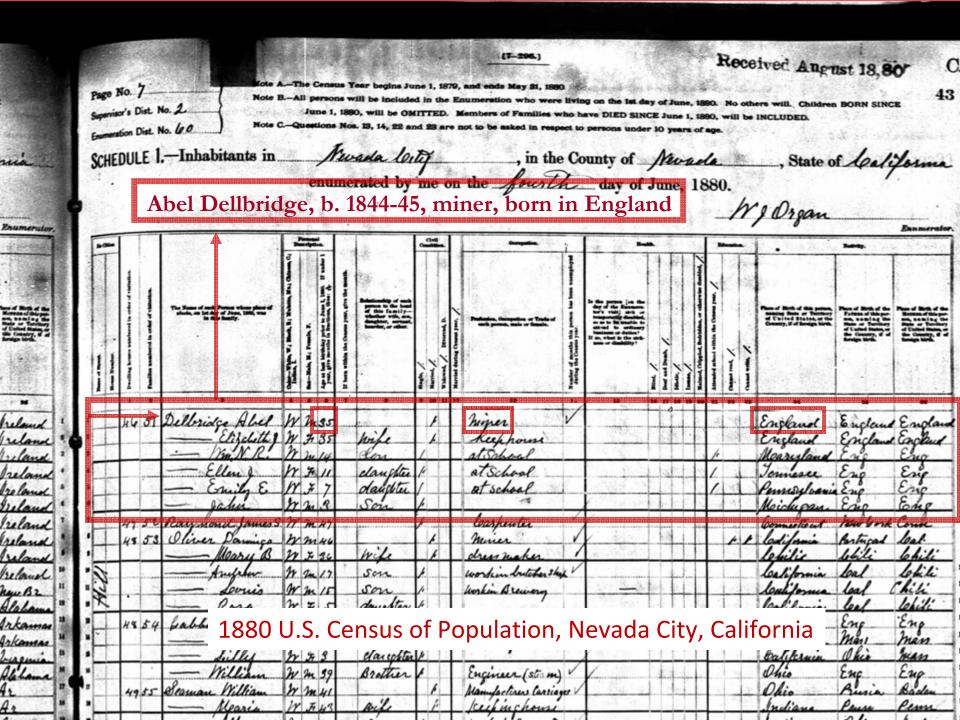
British-born males age 30-39 in the 1880 U.S. Census of Population linked back to the 1861 British Census

British-born males age 30-39 in the 1900 U.S. Census of Population linked back to the 1881 British Census

Main challenge: Lack of specific birthplace info for migrants in U.S. censuses

Requirements/Checks (1880  $\rightarrow$  1861):

- Unique record (name, age birthplace) in 1880 U.S. census and 1861 Br census
- Not present in British 1881 census
- Not present in U.S. 1860 census index
- If they were present in the 1870 U.S. census index, they were not also present in the 1871 British census index, and if they were present in the 1871 British census index, they were not also present in the 1870 U.S. census index.
- Oldest U.S.-born child in 1880 was born after 1860
- Youngest Britain-born child in 1880 was born before 1862



Abel Dellbridge, b. 1844-45, father: miner	1	
Abel Delibridge, b. 1844-45, father: miner	, born in Liskeard, Cornwall, England .	
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U.S. samples: 4,138 (1860-1880) &

3,919 (1880-1900)

British samples: 2,039 (1861-1881) &

4,071 (1881-1901)

Migrant samples: 1,176\* (1861-1880) &

1,144 (1881-1900)

Four occupation categories: White Collar, Farmer, Skilled & Semiskilled, and Unskilled

\* 2,174 linked; remainder awaiting occupational transcription

The conventional approach:

$$\ln Y_i^{\text{Son}} = \beta \ln Y_i^{\text{Father}} + \varepsilon_i$$

where  $\beta$  = "intergenerational income elasticity"

But we've only got occupations, and they're difficult to order unambiguously

9	Britain 1861-1881											
,												
		5	Skilled &									
Son's	White		Semi-	Un-	Row							
Occup.	Collar	Farmer	Skilled	Skilled	Sum							
A. Raw Fre	quencie	S										
(Columr	n Percer	nt)										
WC	117	18	153	54	342							
	(41.9)	(11.6)	(15.8)	(8.5)								
F	3	67	4	10	84							
	(1.1)	(43.2)	(0.4)	(1.6)								
SS	115	46	641	288 1	,090							
	(41.2)	(29.7)	(66.4)	(45.1)								
U	44	24	168	287	523							
	(15.8)	(15.5)	(17.4)	(44.9)								
Col. Sum	279	155	966	639 2	,039							

$$P = \begin{bmatrix} 3 \\ 2 \\ 2 \end{bmatrix} \quad Q = \begin{bmatrix} 2 \\ 4 \\ 6 \end{bmatrix} - \begin{bmatrix} 2 \\ 4 \\ 1 \end{bmatrix} - \begin{bmatrix} 2 \\ 6 \end{bmatrix} \begin{bmatrix} 2 \\ 4 \end{bmatrix} \begin{bmatrix} 2 \\ 4$$

$$P = \begin{bmatrix} 3 & 1 \\ 2 & 2 \end{bmatrix} \quad Q = \begin{bmatrix} 2 & 1 \\ 6 & 1 \end{bmatrix}$$

Cross-Product Ratios:  $(3 \times 2) / (2 \times 1) = 3$  for P  $(2 \times 1) / (6 \times 1) = 1/3$  for Q

$$Q = \begin{bmatrix} 2 & 1 \\ 6 & 1 \end{bmatrix} \qquad Q' = \begin{bmatrix} 2 & 2 \\ 3 & 1 \end{bmatrix}$$

Cross-Product Ratio for Q = ratio for Q' = 1/3

For tables  $> 2 \times 2$ , use the "Altham statistic," which uses all of the cross-product ratios:

$$d(P,Q) = \left[ \sum_{i=1}^{r} \sum_{j=1}^{s} \sum_{l=1}^{r} \sum_{m=1}^{s} \left| ln \left( \frac{p_{ij} p_{lm} q_{im} q_{lj}}{p_{im} p_{lj} q_{ij} q_{lm}} \right) \right|^{2} \right]^{1/2}$$

Measures distance between mobility in P and mobility in Q

#### The Plan:

- For each country/year (e.g. U.S. 1860-80) group occupations into 4 categories (white collar, skilled, farmer, laborer)
- Measure fraction off main diagonal with actual marginal frequencies (M)
- Measure fraction off main diagonal with the marginal frequencies from the other table in the comparison (M')

 Calculate the Altham statistic d(P,J) comparing that 4 × 4 table to independence, a matrix J of ones:

higher  $d(P,J) \Rightarrow$  farther from independence  $\Rightarrow$  less intergenerational mobility

• For country/year pairs (e.g. U.S. 1860-80 & Britain 1861-81) calculate the Altham statistic d(P,Q) to compare the difference in mobility

# Migrants were more mobile at both the top (White Collar) and the bottom (Unskilled)

Father's Occupation

				1 40		<i>-</i> • • • • • • • • • • • • • • • • • • •	ceapadon							
		Brita	in 1861	-1881			Britain 1861-U.S. 1880							
			Skilled &	ķ			Skilled &							
Son's	White		Semi-	Un-	Row		White		Semi-	Un-	Row			
Occup.	Collar	Farme	r Skilled	Skilled	Sum		Collar	Farmer	r Skilled	Skilled	Sum			
A. Raw Fre	equencie	es												
(Colum	n Perce	nt)												
WC	117	18	153	54	342		35	12	113	42	202			
	(41.9)	(11.6)	(15.8)	(8.5)			(24.0)	(13.2)	(16.5)	(16.5)				
F	3	67	4	10	84		15	22	64	49	150			
	(1.1)	(43.2)	(0.4)	(1.6)			(10.3)	(24.2)	(9.3)	(19.3)				
SS	115	46	641	288 1	,090		78	41	439	123	681			
	(41.2)	(29.7)	(66.4)	(45.1)			(53.4)	(45.1)	(64.1)	(48.4)				
U	44	24	168	287	523		18	16	69	40	143			
	(15.8)	(15.5)	(17.4)	(44.9)			(12.3)	(17.6)	(10.1)	(15.7)				
Col. Sum	279	155	966	639 2	2,039		146	91	685	254 1	,176			

Non-migrants ("stayers")

Migrants ("movers")

#### **MOBILITY MEASURES**

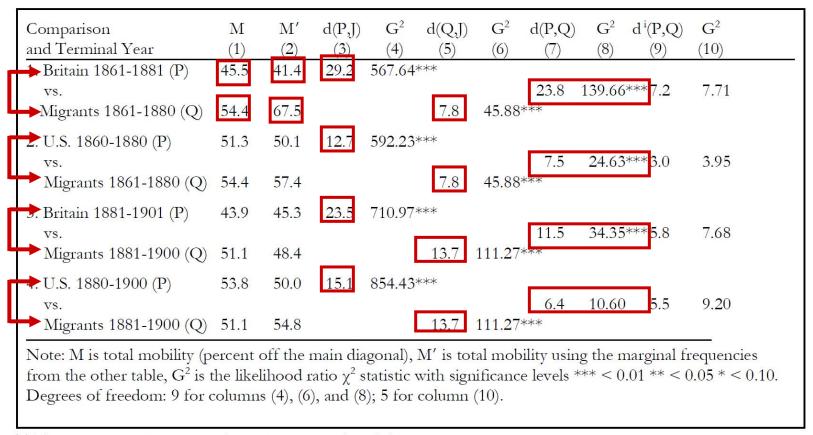
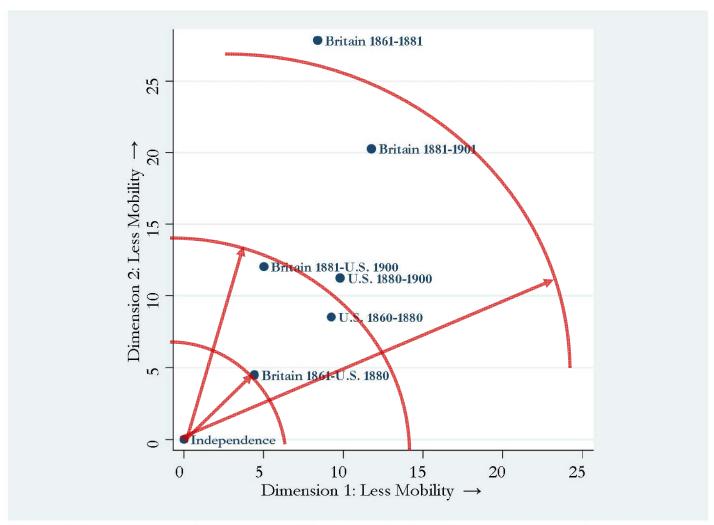


Table 3. Summary Measures of Intergenerational Mobility.

# **VISUALIZING MOBILITY DIFFERENCES**



**Figure 3**. Intergenerational Occupational Mobility in the U.S., Britain, and in British-to-U.S. Migrants (Multidimensional Scaling Scores)

# STRUCTURAL MODEL: SWITCHING ORDERED

We've been *descriptive* up to now, so to move to *causation*, we need to consider selectivity:

$$y_{1i} = \beta_1' X_{1i} + \varepsilon_{1i} \quad \text{if } M_i = 1 \tag{1}$$

$$y_{0i} = \beta_0' X_{0i} + \varepsilon_{0i} \quad \text{if } M_i = 0$$
 (2)

$$M_{i} = \begin{cases} 1 & \text{if } \gamma_{1}'Z_{i} + \gamma_{2}(y_{1i} - y_{0i}) + u_{i} \ge 0\\ 0 & \text{otherwise} \end{cases}$$
(3)

Where y is occupational class, now ordered:

White Collar > Farmer > Skilled & Semiskilled > Unskilled and M=1 if migrant, 0 if non-migrant

#### **SWITCHING ORDERED PROBIT**

# <u>Selection</u> and <u>Treatment Effect</u> parameters:

$$s_{1} = E(y_{1}^{*} | M = 1) - E(y_{1}^{*} | M = 0) = \tilde{X}_{1} \hat{\beta}_{1} - \tilde{X}_{0} \hat{\beta}_{1}$$

$$s_{0} = E(y_{0}^{*} | M = 0) - E(y_{0}^{*} | M = 1) = \tilde{X}_{0} \hat{\beta}_{0} - \tilde{X}_{1} \hat{\beta}_{0}$$

$$(4)$$

$$\tau_1 = E(y_1^* - y_0^* \mid M = 1) = \tilde{X}_1 \hat{\beta}_1 - \tilde{X}_1 \hat{\beta}_0$$
 (5)

$$\tau_0 = E(y_1^* - y_0^* \mid M = 0) = \tilde{X}_0 \hat{\beta}_1 - \tilde{X}_0 \hat{\beta}_0$$
 (6)

#### **SWITCHING ORDERED PROBIT**

O	rdered i	Probit Swit	ching R	egression						
1881	881 <u>M</u> e		Sta	ayers	Structural Probit (Move)					
Characteristic	β	<i>t</i> -stat.	β	<i>t</i> -stat.	β	S.E.	[90%	C.I.]		
Father's Class: 1. WC	0.56	4.39***	0.97	13.35***						
Father's Class: 2. F	0.48	2.30**	0.88	10.02***						
Father's Class: 3. SS	0.25	2.40**	0.37	6.35***						
Age	0.13	0.93	0.10	1.47	0.16	0.15	[-0.10	0.41]		
$Age^2$	0.00	0.85	0.00	1.44	-0.01	0.01	[-0.01	[0.00]		
Father's Age	0.00	0.23	0.01	3.60***	0.00	0.01	[-0.01	0.01]		
Father in Agric	-0.06	0.33	-0.43	6.19***						
One Servant in HH	0.34	2.40**	0.34	4.07***	0.11	0.17	[-0.17	0.40]		
2+ Servants in HH	0.42	2.48**	0.55	4.96***	0.02	0.23	[-0.35	0.40]		
Age Discrepancy	-0.02	0.30	-0.06	2.80***	0.16	0.06	[ 0.06	0.26]		
Eldest Child	-0.06	0.74	-0.04	0.91	-0.12	0.09	[-0.28	0.04]		
Oldest Brother in HI	Η				-0.07	0.06	[-0.17	0.02]		
Children in HH					0.04	0.01	[ 0.02	0.06]		
Mother Employed					-0.19	0.07	[-0.31	-0.07]		
Parish ≠ Birth Parish					-0.04	0.04	[-0.10	0.03]		
$\hat{\mathcal{Y}}_M$ - $\hat{\mathcal{Y}}_S$					-0.90	0.37	[-1.51	-0.29]		
Constant	-0.22	0.21	-0.37	0.74	-2.18	1.08	[-3.95	-0.39]		

Note: Observations: 5,025. Omitted categories are "Father's Class: 4. U," "No Servants in HH," "<2 Servants in HH," "Not Eldest Child," "Not Oldest Brother in HH," "Mother Not

Employed," and "Parish=Birth Parish." Structural Probit SEs and CIs calculated by

bootstrapping via data resampling with 500 repetitions.

Table 4. Ordered Probit Switching Regression (FIML).

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

# **SWITCHING ORDERED PROBIT**

Parameter	Estimate	S.E.	[90%	C.I.]
(1) $\hat{y_M}$ , Movers	1.093	0.127	[ 0.883	1.302]
(2) $\hat{y}_M$ , Stayers	1.028	0.138	[ 0.801	1.254]
(3) $\hat{y}_s$ , Movers	1.128	0.153	[ 0.876	1.380]
(4) $\hat{V}_{\rm s}$ , Stayers	1.002	0.130	[0.787	1.216]
(5) $s_M$ , Selection of migrants=(1)-(2)	0.065	0.027	0.021	0.109
(6) $s_s$ , Selection of stayers=(4)-(3)	-0.127	0.029	[-0.174	-0.080]
(7) $\tau_M$ , Treatment Effect: Treated=(1)-(3)	-0.036	0.196	[-0.359	0.288]
(8) $\tau_s$ , Treatment Effect: Not Treated=(2)-(4)	0.026	0.186	[-0.281	0.333]
Average Treatment Effect	0.013	0.188	[-0.297	0.322]

Table 5. Selection and Treatment Parameters Based On Ordered Probit Switching Regression.

#### **CONCLUSIONS**

Earliest migrants more mobile than both British non-migrants and U.S. native-born

Later migrants still more mobile than British nonmigrants (though gap is smaller) and just as mobile as U.S. native-born

Strong positive selection among migrants

Puzzling result: migration was less likely among those anticipating more improvement

#### **CONCLUSIONS**

#### **Extensions:**

- 1.use country-specific and time-specific occupation incomes instead of categories
- 2.estimate selectivity for first cohort
- 3.examine other outcomes (land ownership) and types of movers (tied vs. independent)

## THE CONTEXT

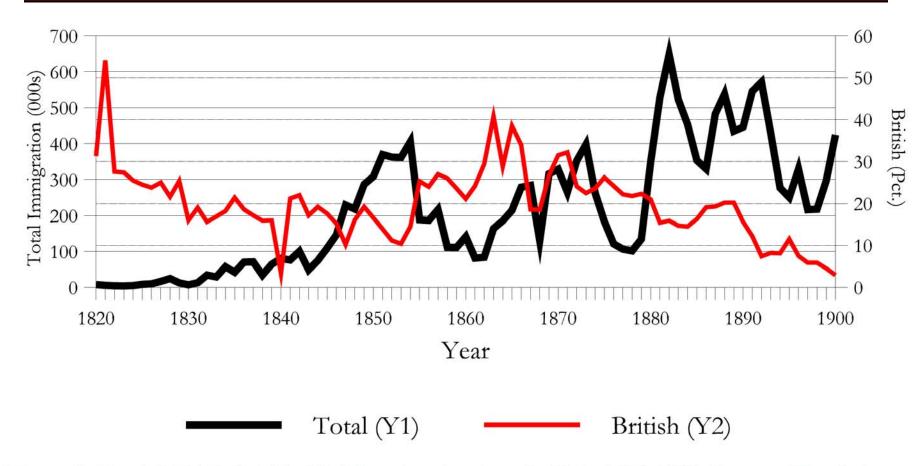
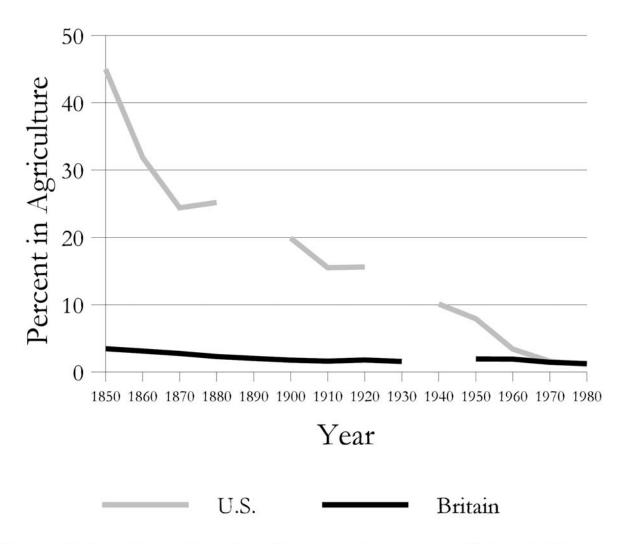


Figure 1. Total (000s) & British (Pct.) Immigration into the U.S., 1820-1900. Source: *Historical Statistics of the U.S. (Millennial Edition)*, Series Ad106-120.

## THE CONTEXT



**Figure 2**. Population Employed in Agriculture in the U.S. and Britain, 1850-1980.

# True whether we look at actual or standardized marginal distributions

Father's Occupation

		Brita	in 1861	-1881		Britain 1861-U.S. 1880							
			Skilled &	ķ		Skilled &							
Son's	White		Semi-	Un-	Row	White		Semi-	Un-	Row			
Occup.	Collar	Farmer	Skilled	Skilled	Sum	Collar	Farmer	Skilled	Skilled	Sum			
B. Standardized Frequencies													
(Colum	n Percei	nt)											
WC	55	6	27	12	100	35	17	27	22	100			
	(55.3)	(5.6)	(27.0)	(12.1)		(34.5)	(16.7)	(26.8)	(22.0)				
F	6	83	3	9	100	17	36	18	30	100			
	(5.7)	(82.7)	(2.8)	(9.0)		(17.1)	(35.5)	(17.6)	(29.7)				
SS	22	6	46	26	100	25	19	34	21	100			
	(22.0)	(5.8)	(45.9)	(26.2)		(25.4)	(18.9)	(34.4)	(21.3)				
U	17	6	24	53	100	23	29	21	27	100			
	(17.0)	(6.1)	(24.2)	(52.7)		(22.9)	(28.8)	(21.2)	(27.1)				
Col. Sum	100	100	100	100	400	100	100	100	100	400			