Cohort Size and The Marriage Market: Explaining a Century of Change in U.S. Marriage Rates

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Cohort Size and The Marriage Market

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• Marriage booms, marriage declines, long-term trends, racial differences

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 - Children's welfare and education
 - Labor force participation and hours of work
 - Income inequality
 - Fertility choices
 - Fraction of individuals on welfare

• In spite of this, there is no general explanation for the variation in marriage formation over time and across races

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Main Contributions

• We provide a general explanation for nearly a century of change in U.S. marriage rates

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 - We first show that one variable, cohort size, explains virtually the entire variation in the U.S. marriage rates since the early 1900s
 - We then develop a simple search model that theoretically can generate the observed pattern in U.S. marriage rates
 - Finally, we test the model and find that it cannot be rejected

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- Existing theories apply only to specific periods
 - Period: Marriage boom after world war II

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- Theories: Labor market opportunities (Wilson, 1987)

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- Theory: Policies (incarceration, welfare aid) (Ellwood and Crane, 1990; Charles and Luoh, 2010)

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 - Period: Early twentieth century, world war I, case studies
 - Theories: Sex Ratio (Becker (1973), Schoen (1983), Guttentag and Secord (1983), Angrist (2002), Abramitzky et al (2011))

• Empirical evidence on marriage rates and cohort size

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- Empirical evidence on marriage rates and cohort size
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- Description of the model
- Tests and results

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Cohort Size and The Marriage Market

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Result on which the paper is based



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Cohort Size and The Marriage Market

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- The mobilization data come from the Selective Service's Special Monograph (1956)

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 - Number of marriages divided by population
 - Share of individuals married within an age range

Empirical Evidence



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- The main issue with this variable is that
 - It is sensitive to changes in population
- If population changes because of migration or other reasons, there are periods in which one draws the wrong inference about the variation in marriage rates

Empirical Evidence



Cohort Size and The Marriage Market

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• This variable conflates two different effects:

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- This variable conflates two different effects:
 - The change in number of people who ever marry
 - The age at first marriage
- If the age at first marriage varies over time, one draws the wrong inference about the variation in marriage rates

• We believe a better measure to capture change in marriage rates is

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 - The share of individuals in a cohort married by a given age
- This variable has two main advantages:
 - It is consistent over time
 - It does not confound different effects

Empirical Evidence



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• Using longitudinal variation, the main result of the paper is that

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- Using longitudinal variation, the main result of the paper is that
 - Cohort size on its own explains almost all of the variation in the marriage rate over time
 - It does this for both blacks and whites



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Cohort Size and The Marriage Market

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• To provide more formal evidence, we regress the log of share ever married on the log of cohort size

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	White Men	Black Men	White Women	Black Women
Ever married	-0.467*	-0.966*	-0.365*	-1.230*
by age 25	(0.107)	(0.102)	(0.075)	(0.113)
R^2	0.21	0.56	0.25	0.63
Ever married	-0.294*	-0.592*	-0.193*	-0.870*
by age 30	(0.037)	(0.048)	(0.026)	(0.064)
R ²	0.50	0.70	0.46	0.74
Ever married	-0.182*	-0.440*	-0.111*	-0.560*
by age 35	(0.015)	(0.031)	(0.011)	(0.043)
R ²	0.71	0.77	0.65	0.74
Ever married	-0.107*	-0.322*	-0.066*	-0.453*
by age 40	(0.008)	(0.019)	(0.007)	(0.026)
R ²	0.77	0.84	0.58	0.85

Time Series Regression of Log Share Ever Married on Log Cohort Size

* Significant at 1%. Standard errors in parentheses.

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Empirical Evidence: Cross-state Variation

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 - If changes in cohort size determine marriage rates, it must be the case that changes in cohort size across states determine changes in marriage rates across states
 - To eliminate possible endogeneity in changes in cohort size due to migration or other variables, we use changes in total births

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- We then compute 10-year differences in the share ever married and regress it on 10-year differences in cohort size

	White	Black	White	Black
	Men	Men	Women	Women
10-Yr. Difference in	-0.080**	-0.090**	-0.064**	-0.092**
Log Cohort Size	(0.028)	(0.024)	(0.017)	(0.019)
N	1 44 ´	`112 ´	1 44	112
R ²	0.42	0.74	0.23	0.53

Cross-Sectional Regression of Difference of Log Share Ever Married by 30

Cross-Sectional Regression of Difference of Log Share Ever Married by 40

	White	Black	White	Black
	Men	Men	Women	Women
10-Yr. Difference in	-0.037**	-0.049*	-0.037**	-0.077**
Log Cohort Size	(0.007)	(0.020)	(0.007)	(0.014)
N	96	74	96	74
R ²	0.30	0.36	0.12	0.51

* Significant at 10%. ** Significant at 1%. Standard errors in parentheses. Time fixed effects.

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- To deal with possible endogeneity issues we use cross-state variation in mobilization rates during WW II to instrument for changes in cohort size (Acemoglu et al (2004))

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- Return effect I: after the war, births could have increased because deployed men who survived the war returned home and started to make up for the missing years
- Return effect II: after the war, births could have decline because many deployed men died

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- We only use white men, ages 18-44
- The Selective Service's Special Monograph provides racial breakdowns of draft registrations and inductions, but not of enlistments
- We construct the mobilization rate measure as the number of white men inducted, divided by the number of white men registered

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The variation in inducted men across states are mainly explained by the following differences:

• Idiosyncratic differences in the behavior of local draft boards

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- Differences in ethnic composition, in particular concentration of Germans and Italians in a state
- Exemptions for farmers: food production was a priority
- Differences in occupational structures: workers employed in industries central to the war effort were exempted

	Low	Medium	High	
Percent Men Inducted into Army	0.28	0.30	0.32**	
	(0.01)	(0.01)	(0.01)	
Share Never Married at 30	0.24	0.22	0.21	
	(0.06)	(0.05)	(0.05)	
Share Farmers	0.24	0.20	0.16	
	(0.13)	(0.09)	(0.14)	
Age	34.84	34.34	34.15*	
	(0.79)	(1.19)	(0.92)	
Men's Employment	0.85	0.84	0.83	
	(0.02)	(0.03)	(0.03)	
Women's Employment	0.26	0.27	0.28	
	(0.08)	(0.05)	(0.07)	
Log Income	6.55	6.59	6.58	
	(0.25)	(0.18)	(0.31)	
Years of Education	9.91	9.61	9.16**	
	(0.62)	(0.70)	(0.53)	
Share German- or Italian-born	0.02	0.02	0.03	
	(0.01)	(0.01)	(0.03)	
Number of Children by Age 35	1.74	2.00	1.90	
	(0.46)	(0.38)	(0.46)	

Cross-state Differences in Main Observable Variables

* Difference between high and low groups significant at 5%. ** Significant at 1%.

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Cohort Size and The Marriage Market

• We consider two specifications

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- In both specifications we control for:
 - State and birth-year fixed effects
 - 1940 education, share farmers, and income interacted with time to allow for potentially different time trends
- In the second specification, we allow the birth-year fixed effect to vary by region

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• In the first stage, we run the following regression:

$$logcohortsize_{c,s} = \sum_{s} \pi_{s} + \sum_{c} \gamma_{c} + \sum_{c} \gamma_{c} \cdot \textit{mobrate}_{s} + X_{c,s}\beta + \nu_{c,s}$$

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- We include the mobilization rates in two ways:
 - The overall mobilization rate in a state interacted with year dummies
 - The previous variables plus the cumulative year-by-year mobilization rates and the war casualty rates by state

First Stage Results. Dependent Variable: Log Conort Size				
	Time FE 1	Time FE 2	Time-Reg. FE 1	Time-Reg. FE 2
Mobilization * 1941	0.227	0.042	0.345	0.291
	(0.483)	(0.505)	(0.545)	(0.545)
Mobilization * 1942	0.681	0.617	1.193**	1.175**
	(0.492)	(0.495)	(0.557)	(0.555)
Mobilization * 1943	-0.151	0.351	0.794	1.098
	(0.500)	(0.694)	(0.568)	(0.733)
Mobilization * 1944	-1.259**	1.528	0.004	1.832
	(0.502)	(1.183)	(0.570)	(1.198)
Mobilization * 1945	-1.545***	-1.584***	-0.606	-0.617
	(0.497)	(0.495)	(0.564)	(0.561)
Mobilization * 1946	-0.835*	-0.874*	-0.061	-0.084
	(0.491)	(0.489)	(0.555)	(0.553)
Mobilization * 1947	-0.604	-0.623	0.033	0.025
	(0.487)	(0.485)	(0.550)	(0.548)
Yr-by-Yr Mobil. * 1943		-0.818		-0.601
		(0.724)		(0.847)
Yr-by-Yr Mobil. * 1944		-2.897***		-1.990*
		(1.111)		(1.150)
Casualty Rate		-2.429		-5.706**
-		(2.408)		(2.428)

* Significant at 10%. ** Significant at 5%. *** Significant at 1%. Standard errors in parentheses.

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Cohort Size and The Marriage Market
• In the second stage, we run the following regression:

$$y_{c,s} = \sum_{s} \pi_{s} + \sum_{c} \gamma_{c} + \phi \cdot \textit{logcohortsize}_{c,s} + X_{c,s}\beta + \epsilon_{c,s}$$

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Second Stage. Regression of Log of Share Ever Married on Log Cohort Size

	No Region FE	No Region FE	Region FE	Region FE
	(1)	(2)	(1)	(2)
(Men)	-0.027*	-0.053*	-0.027*	-0.052*
	(0.003)	(0.01)	(0.003)	(0.01)
(Women)	-0.023*	-0.029*	-0.023*	-0.029*
	(0.001)	(0.01)	(0.001)	(0.01)

* Significant at 1%. Standard errors in parenthesis. Robust standard errors are clustered at the state-level. Each coefficient is the outcome of a separate regression.

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• To understand the mechanism behind the relationship between cohort size and marriage rates we

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 - Develop a simple search model model of the marriage market

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- To understand the mechanism behind the relationship between cohort size and marriage rates we
 - Develop a simple search model model of the marriage market
 - Test it

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• In a search model, to generate a relationship between cohort size and marriage rate, we need the following feature:

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 - The value of search for women declines faster than for men

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- In a search model, to generate a relationship between cohort size and marriage rate, we need the following feature:
 - The value of search for women declines faster than for men
- As a result, women on average marry older men

• In this model, an increase in cohort size has the effect of making older men a scarce resource

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- This change has two main consequences:

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 - First, the fraction of women who marry mechanically declines because fewer of them meet a potential spouse

- In this model, an increase in cohort size has the effect of making older men a scarce resource
- This change has two main consequences:
 - First, the fraction of women who marry mechanically declines because fewer of them meet a potential spouse
 - Second, the fraction of women who marry decreases because young men become more selective

• To capture this we develop a model with the following features:

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- To capture this we develop a model with the following features:
 - $\bullet\,$ The economy is populated by ${\cal T}+1$ overlapping generations of men and women

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- To capture this we develop a model with the following features:
 - The economy is populated by $\mathcal{T}+1$ overlapping generations of men and women
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 - The economy is populated by $\mathcal{T}+1$ overlapping generations of men and women
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 - If in period t an individual of gender i and age a is single, she or he meets a potential spouse with probability θⁱ_{a,t}

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 - $\bullet\,$ The economy is populated by ${\cal T}+1$ overlapping generations of men and women
 - In each period t, a new generation of size N_t is born and lives for T + 1 periods
 - If in period t an individual of gender i and age a is single, she or he meets a potential spouse with probability θⁱ_{a,t}
 - The two spouses then decide whether to marry with the objective of maximizing lifetime utility

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• Women meet a man with a positive probability only in their first period of their life

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- Women meet a man with a positive probability only in their first period of their life
- Men meet a potential spouse with a positive probability in their first two periods of life

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- Men meet a potential spouse with a positive probability in their first two periods of life
- As a result, the marriage market has both young and old men, but only young women

$\bullet\,$ The within-period utility of being single is $\delta\,$

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- ${\, \bullet \,}$ The within-period utility of being single is δ
- The within-period utility of being married for the couple will be denoted by η , which does not change over time

- $\bullet\,$ The within-period utility of being single is $\delta\,$
- The within-period utility of being married for the couple will be denoted by η , which does not change over time
- η is drawn from a distribution $F(\eta)$ which does not depend on the age of the couple

 $\bullet\,$ The utility from future periods is discounted at the discount factor $\beta \leq 1\,$

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- The couple's lifetime utility is divided between spouse *i* and spouse *j* using a Nash bargaining solution:

$$w_t^i(\eta) = v_t^i + \gamma_i \left[v_t - v_t^i - v_t^j \right]$$

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$$w_t^i(\eta) = v_t^i + \gamma_i \left[v_t - v_t^i - v_t^j \right]$$

• There is no divorce

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• Hence, they marry if
$$\eta' > \underline{\eta}_{old} = 2\delta$$
.

The model

• If a woman meets a young man, they marry if they draw a marriage utility η that is greater than

$$\underline{\eta}_{young} = 2\delta + \beta \frac{1 - \beta^{T}}{1 - \beta^{T+1}} \gamma \left\{ E\left[\eta \left| \eta \ge 2\delta\right] - 2\delta \right\} \left(1 - F\left(2\delta\right)\right) \theta_{1,t}^{m} \right\}$$

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 Where θ^m_{1,t} is the probability that an young man meets a woman next period:

$$\theta_{1,t}^{m} = \frac{N_{0,t}}{N_{0,t} + N_{1,t}^{m}} = \theta_{0,t}^{m}$$

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• Notice that in the model cohort size has a direct effect only through the matching probabilities and the reservation value of young men

• Using the model, we show the following result:

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Proposition

A positive and permanent shock to cohort size in period τ reduces the fraction of cohort τ individuals who get married. A negative and permanent shock in period τ has the opposite effect.
• Using the model, we show the following result:

Proposition

A positive and permanent shock to cohort size in period τ reduces the fraction of cohort τ individuals who get married. A negative and permanent shock in period τ has the opposite effect.

• This Proposition shows that in a search model a change in cohort size has the desired effect on marriage rates

• The following Proposition describes the implication we will use to test the search model:

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• The following Proposition describes the implication we will use to test the search model:

Proposition

In the search model we consider, an increase in cohort size reduces the average age difference between spouses. A reduction in cohort size has the opposite effect.

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• We will first show graphically the relationship between cohort size and average age difference by race

- We will first show graphically the relationship between cohort size and average age difference by race
- We will then estimate the effect of cohort size on average age differences

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Outcome of the Test: Whites

Figure: Age Difference Between Spouses by Cohort, Whites



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Regressions: Log Age Difference and Log Cohort Size

	Time Series,	Cross-State,	Cross-State,	Cross-State,
	1930-1975	1940-1950	1950-1960	1960-1970
Log Cohort Size	-0.592***			
	(0.043)			
10-Yr. Difference in		0.110	-0.172**	-0.130*
Log Cohort Size		(0.235)	(0.077)	(0.078)
N	46	49	49	49
R-squared	0.81	0.00	0.05	0.09

* Significant at 10%. ** Significant at 5%. *** Significant at 1%.

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• Given that the model is not rejected, we then structurally estimate it

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- The goal is to determine whether quantitatively we can match the patterns observed in the data

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- Given that the model is not rejected, we then structurally estimate it
- The goal is to determine whether quantitatively we can match the patterns observed in the data
- The model is estimated using Simulated Method of Moments

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• In the estimation we make the following additional assumptions:

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- In the estimation we make the following additional assumptions:
 - The value of marriage η is drawn from a beta distribution with shape parameters α_1 and α_2

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- In the estimation we make the following additional assumptions:
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$$\delta_t = \delta + \nu_t$$

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 - The value of marriage η is drawn from a beta distribution with shape parameters α_1 and α_2
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• The annual discount factor is set equal to 0.98

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The Model: Estimation

• (Cont'd) Assumptions:

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- (Cont'd) Assumptions:
 - Each period in the model corresponds to 10 years

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 - Each period in the model corresponds to 10 years
 - Each generation starts making decisions at age 20 and lives for 50 years (5 periods)

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 - Each period in the model corresponds to 10 years
 - Each generation starts making decisions at age 20 and lives for 50 years (5 periods)
 - $\bullet\,$ We assume that Nash-bargaining is symmetric by setting $\gamma\,$ equal to 0.5

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- (Cont'd) Assumptions:
 - Each period in the model corresponds to 10 years
 - Each generation starts making decisions at age 20 and lives for 50 years (5 periods)
 - $\bullet\,$ We assume that Nash-bargaining is symmetric by setting $\gamma\,$ equal to 0.5
 - $\bullet\,$ We augment the model to allow for a fraction of men that are unwilling to marry no matter the value of match quality $1-\phi$

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• We therefore estimate the following 4 parameters:

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 - The constant in the value of being single δ which is assumed to be identical across gender and over time

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- We therefore estimate the following 4 parameters:
 - The constant in the value of being single δ which is assumed to be identical across gender and over time
 - The two shape parameters α_1 and α_2
 - $\bullet\,$ The fraction of men unwilling to marry $\phi\,$

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• We match 51 moments:

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- We match 51 moments:
 - The fraction of women never married in a cohort starting from the cohort born in 1930 and ending with the cohort born in 1980

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- We match 51 moments:
 - The fraction of women never married in a cohort starting from the cohort born in 1930 and ending with the cohort born in 1980
- The goal of the exercise is to understand whether the model can quantitatively match the patterns observed in the data

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Estimated Parameters

Parameters	Estimates	Standard Errors
First Shape Parameter	0.020	[0.010]
Second Shape Parameter	0.071	[0.039]
Value of Being Single	0.107	[0.037]
Fraction of Men Unwilling to Marry	13.3	[0.143]

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The Model: Estimation



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• We show that cohort size on its own explains a large fraction of the variation in the marriage rates over the last century

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- We then develop a simple search model that has the potential of generating the relationship between changes in cohort size and changes in marriage rates

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- We show that cohort size on its own explains a large fraction of the variation in the marriage rates over the last century
- We then develop a simple search model that has the potential of generating the relationship between changes in cohort size and changes in marriage rates
- We test the model and show that it is consistent with the patterns observed in the data

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THE END

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Empirical Evidence



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